# universität innsbruck



# **Hidden Markov Models**

703075. Machine Learning https://iis.uibk.ac.at/courses/2020s/703075

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#### Hidden Markov models

Introduction Discrete Markov process Observable Markov model Hidden Markov model (HMM) Solving HMMs

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## Introduction

- Modeling dependencies in input; no longer iid
- Sequences:
  - Temporal: In speech; phonemes in a word (dictionary), words in a sentence (syntax, semantics of the language).
    - In handwriting, pen movements
  - Spatial: In a DNA sequence; base pairs

## Introduction

- Modeling dependencies in input; no longer iid
- Sequences:
  - A sequence can be characterized as being generated by a *parametric random process*

#### Hidden Markov models

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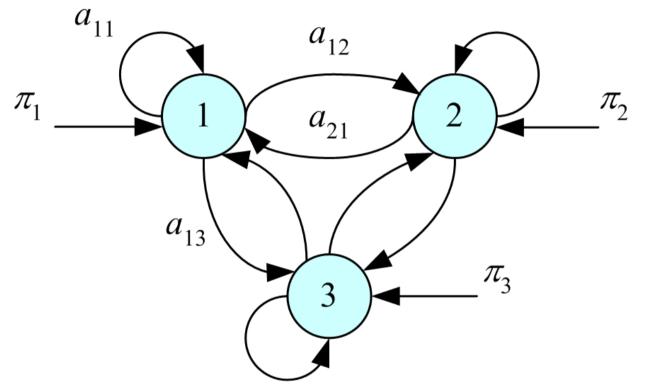
Transition probabilities

$$a_{ij} \equiv P(q_{t+1}=S_j \mid q_t=S_i)$$
  $a_{ij} \ge 0 \text{ and } \sum_{j=1}^N a_{ij}=1$ 

Initial probabilities

$$\pi_i \equiv P(q_1 = S_i) \qquad \sum_{j=1}^N \pi_i = 1$$

#### **Stochastic Automation**



• Transition probabilities

$$a_{ij} = P(q_{t+1} = S_j | q_t = S_i)$$
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#### **Observable Markov model**

• The states are observable

- At any time *t* we know  $q_t$ 

Having an observation sequence

$$O = Q = \{q_1 q_2 \dots q_T\}$$
  

$$P(O = Q | \mathbf{A}, \mathbf{\Pi}) = P(q_1) \prod_{t=2}^{T} P(q_t | q_{t-1}) = \pi_{q_1} a_{q_1 q_2} \cdots a_{q_{T-1} q_T}$$

- Rabiner and Juang (1986)
- Three urns each full of balls of one color

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$$S_1: \text{ red, } S_2: \text{ blue, } S_3: \text{ green}$$
$$\Pi = \begin{bmatrix} 0.5, 0.2, 0.3 \end{bmatrix} \quad \mathbf{A} = \begin{bmatrix} 0.4 & 0.3 & 0.3 \\ 0.2 & 0.6 & 0.2 \\ 0.1 & 0.1 & 0.8 \end{bmatrix}$$

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• Learning

- Given K example sequences of length T

$$\hat{\pi}_{i} = \frac{\#\{\text{sequences starting with } S_{i}\}}{\#\{\text{sequences}\}} = \frac{\sum_{k} l(q_{1}^{k} = S_{i})}{K}$$

• Learning

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$$\begin{aligned} \hat{\pi}_{i} &= \frac{\#\{\text{sequences starting with } S_{i}\}}{\#\{\text{sequences}\}} = \frac{\sum_{k} l(q_{1}^{k} = S_{i})}{K} \\ \hat{a}_{ij} &= \frac{\#\{\text{transitions from } S_{i} \text{ to } S_{j}\}}{\#\{\text{transitions from } S_{i}\}} \\ &= \frac{\sum_{k} \sum_{t=1}^{T-1} l(q_{t}^{k} = S_{i} \text{ and } q_{t+1}^{k} = S_{j})}{\sum_{k} \sum_{t=1}^{T-1} l(q_{t}^{k} = S_{i})} \end{aligned}$$

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## Hidden Markov Models

- States are not observable
- Discrete observations {v<sub>1</sub>,v<sub>2</sub>,...,v<sub>M</sub>} are recorded
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  - A probabilistic function of the state
- Emission probabilities

- Observation that we observe

 $v_m$ ,  $m = 1, \ldots, M$  in state  $S_i$ 

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 $v_m$ ,  $m = 1, \ldots, M$  in state  $S_j$ 

$$b_j(m) \equiv P(O_t = v_m | q_t = S_j)$$

The state sequence Q is not observed

 but it should be inferred from the observation sequence O

- In each urn, there are balls of different colors, but with different probabilities.
  - For each observation sequence, there are multiple state sequences

 $-b_j(m) = P(O_t = v_m | q_t = S_j) \text{ denotes the probability of}$ drawing a ball of color *m* from urn *j* 

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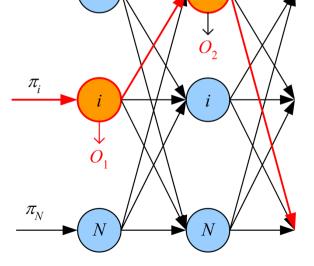
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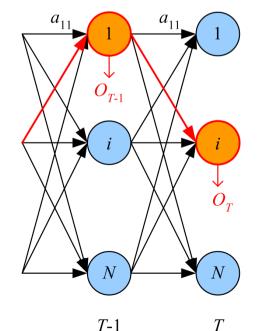
 We again observe a sequence of ball colors but without knowing the sequence of urns from which the balls were drawn

- In each urn, there are balls of different colors, but with different probabilities.
  - The observable model is a special case of the hidden model
  - where M = N
  - and b(m) is 1 if j =m and 0 otherwise

- In each urn, there are balls of different colors, but with different probabilities.
  - For the same observation sequence O, there may be many possible state sequences Q that could

have generated  $O_{\frac{\pi_1}{\mu}}$ 





*T*-1

#### **Elements of an HMM**

- *N*: Number of states  $S = \{S_1, S_2, ..., S_N\}$
- *M*: Number of observation symbols

$$V = \{v_1, v_2, \ldots, v_M\}$$

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- *N*: Number of states  $S = \{S_1, S_2, ..., S_N\}$
- *M*: Number of observation symbols  $V = \{v_1, v_2, \dots, v_M\}$
- $\mathbf{A} = [a_{ij}]: N \text{ by } N \text{ state transition probability matrix}$  $\mathbf{A} = [a_{ij}] \text{ where } a_{ij} \equiv P(q_{t+1} = S_j | q_t = S_i)$
- $\mathbf{B} = b_j(m)$ : *N* by *M* observation probability matrix  $\mathbf{B} = [b_j(m)]$  where  $b_j(m) \equiv P(O_t = v_m | q_t = S_j)$
- $\Pi = [\pi_i]$ : *N* by 1 initial state probability vector  $\Pi = [\pi_i]$  where  $\pi_i \equiv P(q_1 = S_i)$

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 $\lambda = (\mathbf{A}, \mathbf{B}, \mathbf{\Pi})$ , parameter set of HMM

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# Three basic problems of HMMs

- Given a number of sequences of observations, we are interested in:
  - 1. Evaluation

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3. Learning

Given  $X = \{O^k\}_k$ , find  $\lambda^*$  such that

 $P(X | \lambda^*) = \max_{\lambda} P(X | \lambda)$ 

#### 1. Evaluation

• Given an observation sequence  $O = \{O_1 \ O_2 \cdots O_T\}$ and a state sequence  $Q = \{q_1q_2 \cdots q_T\}$ , the probability of observing O given the state sequence Qis

$$P(O|Q,\lambda) = \prod_{t=1}^{I} P(O_t|q_t,\lambda) = b_{q_1}(O_1) \cdot b_{q_2}(O_2) \cdots b_{q_T}(O_T)$$

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$$P(O|Q,\lambda) = \prod_{t=1}^{T} P(O_t|q_t,\lambda) = b_{q_1}(O_1) \cdot b_{q_2}(O_2) \cdots b_{q_T}(O_T)$$

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- But, we can compute the probability of a state sequence  $P(Q|\lambda) = P(q_1) \prod_{t=2} P(q_t|q_{t-1}) = \pi_{q_1} a_{q_1q_2} \cdots a_{q_{T-1}q_T}$

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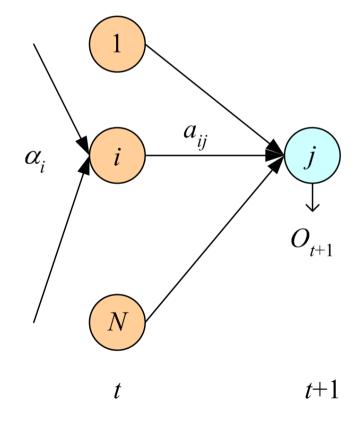
- Which we cannot calculate because we do not know the state sequence.
- But, we can compute

$$P(O|\lambda) = \sum_{\text{all possible } Q} P(O, Q|\lambda)$$

• There are  $N^T$  possible Q

- Forward-Backward procedure
  - Forward variable

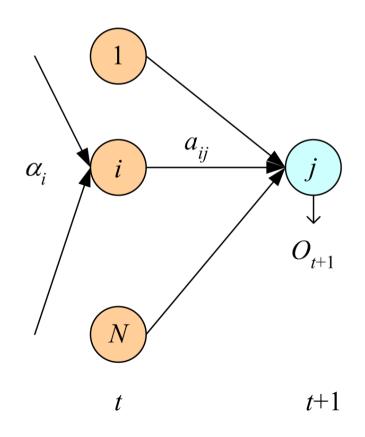
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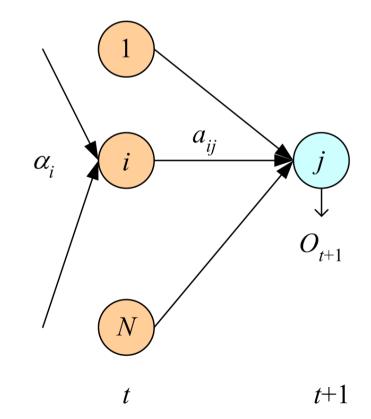


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$$\alpha_1(i) \equiv P(O_1, q_1 = S_i | \lambda)$$
  
=  $P(O_1 | q_1 = S_i, \lambda) P(q_1 = S_i | \lambda)$   
=  $\pi_i b_i(O_1)$ 

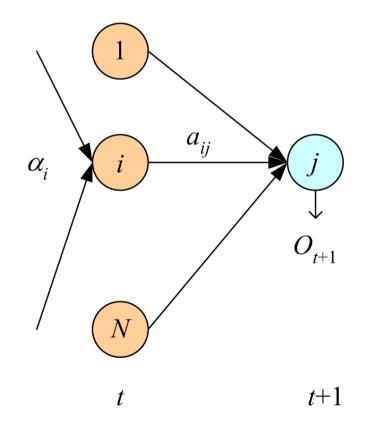


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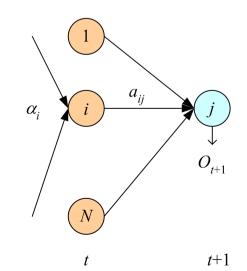
$$\alpha_1(i) = \pi_i b_i(O_1)$$

$$\alpha_{t+1}(j) = \left[\sum_{i=1}^{N} \alpha_t(i) a_{ij}\right] b_j(O_{t+1})$$

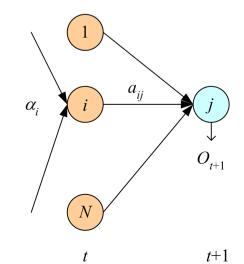


**1. Evaluation**  
$$P(O|\lambda) = \sum_{i=1}^{N} P(O, q_T = S_i | \lambda)$$

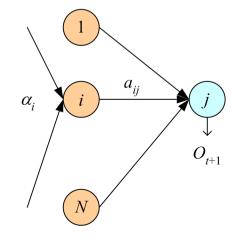
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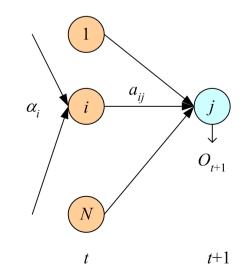
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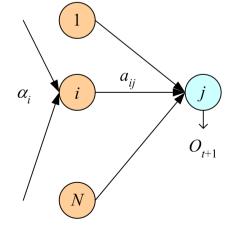


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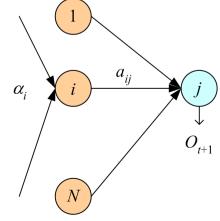
#### Recursion

$$\begin{aligned} \alpha_{t+1}(j) &\equiv P(O_1 \cdots O_{t+1}, q_{t+1} = S_j | \lambda) \\ &= P(O_1 \cdots O_{t+1} | q_{t+1} = S_j, \lambda) P(q_{t+1} = S_j | \lambda) \\ &= P(O_1 \cdots O_t | q_{t+1} = S_j, \lambda) P(O_{t+1} | q_{t+1} = S_j, \lambda) P(q_{t+1} = S_j | \lambda) \\ &= P(O_1 \cdots O_t, q_{t+1} = S_j | \lambda) P(O_{t+1} | q_{t+1} = S_j, \lambda) \\ &= P(O_{t+1} | q_{t+1} = S_j, \lambda) \sum_i P(O_1 \cdots O_t, q_t = S_i, q_{t+1} = S_j | \lambda) \end{aligned}$$



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#### Recursion

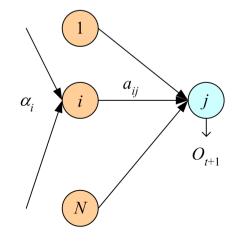
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$$= P(O_{t+1}|q_{t+1} = S_j, \lambda)$$

$$\sum_{i} P(O_1 \cdots O_t, q_{t+1} = S_j|q_t = S_i, \lambda) P(q_t = S_i|\lambda)$$

$$= P(O_1 \cdots P_t, q_{t+1} = S_t, \lambda)$$

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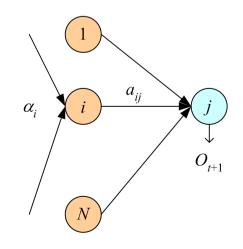
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$$\sum_{i} P(O_1 \cdots O_t, q_{t+1} = S_j|q_t = S_i, \lambda) P(q_t = S_i|\lambda)$$

$$= P(O_{t+1}|q_{t+1} = S_j, \lambda)$$

$$\sum_{i} P(O_1 \cdots O_t | q_t = S_i, \lambda) P(q_{t+1} = S_j | q_t = S_i, \lambda) P(q_t = S_i | \lambda)$$

$$= P(O_{t+1}|q_{t+1} = S_j, \lambda)$$
  
$$\sum_i P(O_1 \cdots O_t, q_t = S_i|\lambda) P(q_{t+1} = S_j|q_t = S_i, \lambda)$$



#### Recursion

$$\alpha_{t+1}(j) \equiv P(O_1 \cdots O_{t+1}, q_{t+1} = S_j | \lambda)$$

$$= P(O_{t+1}|q_{t+1} = S_j, \lambda)$$

$$\sum_{i} P(O_1 \cdots O_t, q_{t+1} = S_j|q_t = S_i, \lambda) P(q_t = S_i|\lambda)$$

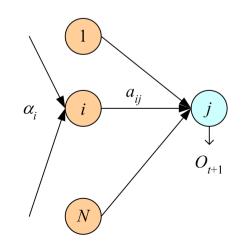
$$= P(O_{t+1}|q_{t+1} = S_j, \lambda)$$

$$\sum_{i} P(O_1 \cdots O_t | q_t = S_i, \lambda) P(q_{t+1} = S_j | q_t = S_i, \lambda) P(q_t = S_i | \lambda)$$

$$= P(O_{t+1}|q_{t+1} = S_j, \lambda)$$

$$\sum_i P(O_1 \cdots O_t, q_t = S_i|\lambda) P(q_{t+1} = S_j|q_t = S_i, \lambda)$$

$$= \left[\sum_{i=1}^N \alpha_t(i) a_{ij}\right] b_j(O_{t+1})$$

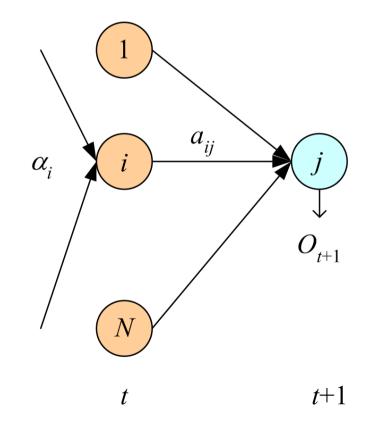


- Forward-Backward procedure
  - Forward variable

$$\alpha_t(i) \equiv P(O_1 \cdots O_t, q_t = S_i | \lambda)$$
  
Initialization:

$$\alpha_1(i) = \pi_i b_i(O_1)$$

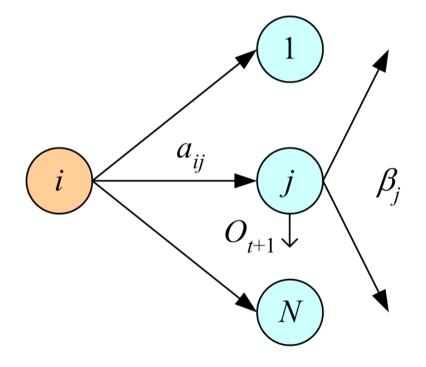
$$\alpha_{t+1}(j) = \left[\sum_{i=1}^{N} \alpha_t(i) a_{ij}\right] b_j(O_{t+1})$$
$$P(O \mid \lambda) = \sum_{i=1}^{N} \alpha_T(i)$$



$$P(O|Q,\lambda) = \prod_{t=1}^{T} P(O_t|q_t,\lambda) = b_{q_1}(O_1) \cdot b_{q_2}(O_2) \cdots b_{q_T}(O_T)$$

- Forward-Backward procedure
  - Backward variable

$$\beta_t(i) \equiv P(O_{t+1} \cdots O_T \mid q_t = S_i, \lambda)$$



t

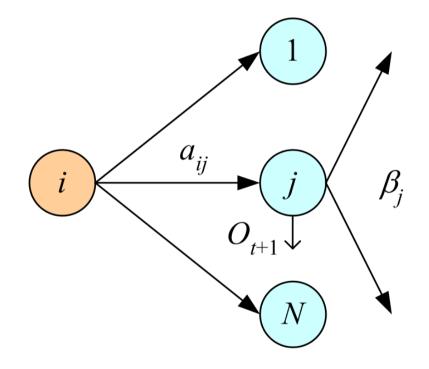
$$P(O|Q,\lambda) = \prod_{t=1}^{T} P(O_t|q_t,\lambda) = b_{q_1}(O_1) \cdot b_{q_2}(O_2) \cdots b_{q_T}(O_T)$$

- Forward-Backward procedure
  - Backward variable

$$\beta_t(i) \equiv P(O_{t+1} \cdots O_T \mid q_t = S_i, \lambda)$$

Initialization:

$$\beta_{\tau}(i) = 1$$



t

$$P(O|Q,\lambda) = \prod_{t=1}^{T} P(O_t|q_t,\lambda) = b_{q_1}(O_1) \cdot b_{q_2}(O_2) \cdots b_{q_T}(O_T)$$

- Forward-Backward procedure
  - Backward variable

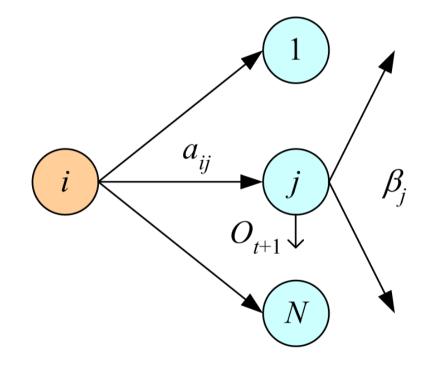
$$\beta_t(i) \equiv P(O_{t+1} \cdots O_T \mid q_t = S_i, \lambda)$$

Initialization:

$$\beta_{\tau}(i) = 1$$

Recursion:

$$\beta_t(i) = \sum_{j=1}^N a_{ij} b_j(O_{t+1}) \beta_{t+1}(j)$$



t

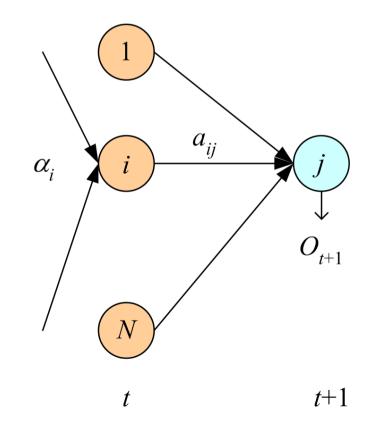
$$P(O|Q,\lambda) = \prod_{t=1}^{T} P(O_t|q_t,\lambda) = b_{q_1}(O_1) \cdot b_{q_2}(O_2) \cdots b_{q_T}(O_T)$$

- Forward-Backward procedure
  - Forward variable

$$\alpha_t(i) \equiv P(O_1 \cdots O_t, q_t = S_i | \lambda)$$
  
Initialization:

$$\alpha_1(i) = \pi_i b_i(O_1)$$

$$\alpha_{t+1}(j) = \left[\sum_{i=1}^{N} \alpha_t(i) a_{ij}\right] b_j(O_{t+1})$$
$$P(O \mid \lambda) = \sum_{i=1}^{N} \alpha_T(i)$$



• Find the state sequence  $Q = \{q_1q_2 \cdots q_T\}$  having the highest probability of generating the observation sequence  $O = \{O_1 \ O_2 \cdots O_T\}$ , given the model  $\lambda$ 

$$P(Q^* \mid O, \lambda) = \max_Q P(Q \mid O, \lambda)$$

$$P(Q^* | O, \lambda) = \max_Q P(Q | O, \lambda)$$

• Let us define  $\gamma_t(i)$  as the probability of being in state  $S_i$  at time t, given O and  $\lambda$   $\gamma_t(i) \equiv P(q_t = S_i | O, \lambda)$  $= \frac{P(O|q_t = S_i, \lambda)P(q_t = S_i|\lambda)}{P(O|\lambda)}$ 

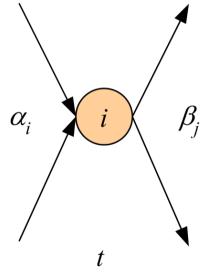
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• Let us define  $\gamma_t(i)$  as the probability of being in state  $S_i$  at time t, given O and  $\lambda$ 

$$\gamma_{t}(i) \equiv P(q_{t} = S_{i} | O, \lambda)$$
$$= \frac{\alpha_{t}(i)\beta_{t}(i)}{\sum_{j=1}^{N} \alpha_{t}(j)\beta_{t}(j)}$$



- Choose the state that has the highest probability
- for each time step:

$$q_t^* = \arg \max_i \gamma_t(i)$$

# 2. State sequence $\gamma_{t}(i) \equiv P(q_{t} = S_{i} | O, \lambda)$ $= \frac{\alpha_{t}(i)\beta_{t}(i)}{\sum_{j=1}^{N} \alpha_{t}(j)\beta_{t}(j)}$

- Choose the state that has the highest probability
- for each time step:  $q_t^* = \arg \max_i \gamma_t(i)$ . NO
- Viterbi Algorithm

2. State sequence  $\gamma_{t}(i) \equiv P(q_{t} = S_{i} | O, \lambda)$   $= \frac{\alpha_{t}(i)\beta_{t}(i)}{\sum_{j=1}^{N} \alpha_{t}(j)\beta_{t}(j)}$ 

• Viterbi Algorithm

- Given state sequence  $Q = q_1 q_2 \cdots q_T$  and observation sequence  $O = O_1 \cdots O_T$ , we define  $\delta_t(i)$  as the probability of the highest probability path at time *t* that accounts for the first *t* observations and ends in  $S_i$ 

 $\delta_t(i) \equiv \max_{q_1 q_2 \cdots q_{t-1}} p(q_1 q_2 \cdots q_{t-1}, q_t = S_i, O_1 \cdots O_t | \lambda)$ 

 $\delta_t(i) \equiv \max_{q_1q_2\cdots q_{t-1}} p(q_1q_2\cdots q_{t-1}, q_t = S_i, O_1\cdots O_t | \lambda)$ 

- Viterbi Algorithm
  - Initialization:

 $\delta_1(i) = \pi_i b_i(O_1)$  $\psi_1(i) = 0$ 

 $\delta_t(i) \equiv \max_{q_1 q_2 \cdots q_{t-1}} p(q_1 q_2 \cdots q_{t-1}, q_t = S_i, O_1 \cdots O_t | \lambda)$ 

- Viterbi Algorithm
  - Initialization:

 $\delta_1(i) = \pi_i b_i(O_1)$  $\psi_1(i) = 0$ - Recursion

$$\delta_t(j) = \max_i \delta_{t-1}(i) a_{ij} b_j(O_t)$$
$$\psi_t(j) = \operatorname{argmax}_i \delta_{t-1}(i) a_{ij}$$

 $\delta_t(i) \equiv \max_{q_1 q_2 \cdots q_{t-1}} p(q_1 q_2 \cdots q_{t-1}, q_t = S_i, O_1 \cdots O_t | \lambda)$ 

- Viterbi Algorithm
  - Initialization:

 $\delta_1(i) = \pi_i b_i(O_1)$  $\psi_1(i) = 0$ - Recursion

-Termination:

$$p^* = \max_i \delta_T(i)$$
$$q_T^* = \operatorname{argmax}_i \delta_T(i)$$

 $\delta_t(j) = \max_i \delta_{t-1}(i) a_{ij} b_j(O_t)$  $\psi_t(j) = \operatorname{argmax}_i \delta_{t-1}(i) a_{ij}$ 

 $\delta_t(i) \equiv \max_{q_1 q_2 \cdots q_{t-1}} p(q_1 q_2 \cdots q_{t-1}, q_t = S_i, O_1 \cdots O_t | \lambda)$ 

- Viterbi Algorithm
  - Initialization:

-Termination:

 $\delta_{1}(i) = \pi_{i}b_{i}(O_{1}) \qquad p^{*} = \max_{i}\delta_{T}(i)$   $\psi_{1}(i) = 0 \qquad q_{T}^{*} = \operatorname{argmax}_{i}\delta_{T}(i)$ - Recursion

$$\delta_t(j) = \max_i \delta_{t-1}(i)a_{ij}b_j(O_t)$$
  
$$\psi_t(j) = \operatorname{argmax}_i \delta_{t-1}(i)a_{ij}$$
  
- Path bactracking

$$q_t^* = \psi_{t+1}(q_{t+1}^*), t=T-1, T-2, ..., 1$$

 $\delta_t(i) \equiv \max_{q_1 q_2 \cdots q_{t-1}} p(q_1 q_2 \cdots q_{t-1}, q_t = S_i, O_1 \cdots O_t | \lambda)$ 

• Viterbi Algorithm

– Recursion

– Termination:

 $p^* = \max_i \delta_T(i)$  $q_T^* = \operatorname{argmax}_i \delta_T(i)$ 

 $\psi_t(j) = \operatorname{argmax}_i \delta_{t-1}(i) a_{ii}$ 

 $\delta_t(j) = \max_i \delta_{t-1}(i) a_{ij} b_i(O_t)$ 

Path bactracking

$$q_t^* = \psi_{t+1}(q_{t+1}^*), t=T-1, T-2, ..., 1$$

 $-\psi_t(j)$  keeps track of the state that maximizes  $\delta_t(j)$  at time t-1, that is, the best previous state

## 3. Learning

Maximum Likelihood

- Calculate  $\lambda *$  that maximizes the likelihood of the sample of training sequences,

$$X = \{O^k\}_{k=1}^K$$
, namely,  $P(X|\lambda)$ 

- Maximum Likelihood
  - Calculate  $\lambda *$  that maximizes the likelihood of the sample of training sequences,

$$\mathbf{X} = \{O^k\}_{k=1}^{K}$$
, namely,  $P(\mathbf{X}|\lambda)$ 

$$\xi_t(i,j) \equiv P(q_t = S_i, q_{t+1} = S_j \mid O, \lambda)$$

$$\begin{aligned} \xi_t(i,j) &= P(q_t = S_i, q_{t+1} = S_j \mid O, \lambda) \\ \xi_t(i,j) &= P(q_t = S_i, q_{t+1} = S_j \mid O, \lambda) \\ &= \frac{P(O \mid q_t = S_i, q_{t+1} = S_j, \lambda) P(q_t = S_i, q_{t+1} = S_j \mid \lambda)}{P(O \mid \lambda)} \end{aligned}$$

$$\begin{split} \xi_{t}(i,j) &= P(q_{t} = S_{i}, q_{t+1} = S_{j} \mid O, \lambda) \\ \xi_{t}(i,j) &= P(q_{t} = S_{i}, q_{t+1} = S_{j} \mid O, \lambda) \\ &= \frac{P(O \mid q_{t} = S_{i}, q_{t+1} = S_{j}, \lambda) P(q_{t} = S_{i}, q_{t+1} = S_{j} \mid \lambda)}{P(O \mid \lambda)} \\ &= \frac{P(O \mid q_{t} = S_{i}, q_{t+1} = S_{j}, \lambda) P(q_{t+1} = S_{j} \mid q_{t} = S_{i}, \lambda) P(q_{t} = S_{i} \mid \lambda)}{P(O \mid \lambda)} \\ &= \left(\frac{1}{P(O \mid \lambda)}\right) P(O_{1} \cdots O_{t} \mid q_{t} = S_{i}, \lambda) P(O_{t+1} \mid q_{t+1} = S_{j}, \lambda) \\ P(O_{t+2} \cdots O_{T} \mid q_{t+1} = S_{j}, \lambda) a_{ij} P(q_{t} = S_{i} \mid \lambda) \end{split}$$

$$\begin{split} \xi_t(i,j) &= P(q_t = S_i, q_{t+1} = S_j | O, \lambda) \\ \xi_t(i,j) &= P(q_t = S_i, q_{t+1} = S_j | O, \lambda) \\ &= \frac{P(O|q_t = S_i, q_{t+1} = S_j, \lambda) P(q_t = S_i, q_{t+1} = S_j | \lambda)}{P(O|\lambda)} \\ &= \frac{P(O|q_t = S_i, q_{t+1} = S_j, \lambda) P(q_{t+1} = S_j | q_t = S_i, \lambda) P(q_t = S_i | \lambda)}{P(O|\lambda)} \\ &= \left(\frac{1}{P(O|\lambda)}\right) P(O_1 \cdots O_t | q_t = S_i, \lambda) P(O_{t+1} | q_{t+1} = S_j, \lambda) \\ P(O_{t+2} \cdots O_T | q_{t+1} = S_j, \lambda) a_{ij} P(q_t = S_i | \lambda) \\ &= \left(\frac{1}{P(O|\lambda)}\right) P(O_1 \cdots O_t, q_t = S_i | \lambda) P(O_{t+1} | q_{t+1} = S_j, \lambda) \\ P(O_{t+2} \cdots O_T | q_{t+1} = S_j, \lambda) a_{ij} \\ &= \frac{\alpha_t(i) b_j(O_{t+1}) \beta_{t+1}(j) a_{ij}}{\sum_k \sum_l P(q_t = S_k, q_{t+1} = S_l, O| \lambda)} \end{split}$$

$$\begin{split} \xi_t(i,j) &= P(q_t = S_i, q_{t+1} = S_j \mid O, \lambda) \\ \xi_t(i,j) &= P(q_t = S_i, q_{t+1} = S_j \mid O, \lambda) \\ &= \frac{P(O\mid q_t = S_i, q_{t+1} = S_j, \lambda) P(q_t = S_i, q_{t+1} = S_j \mid \lambda)}{P(O\mid \lambda)} \\ &= \left(\frac{1}{P(O\mid\lambda)}\right) P(O_1 \cdots O_t, q_t = S_i \mid \lambda) P(O_{t+1} \mid q_{t+1} = S_j, \lambda) \\ P(O_{t+2} \cdots O_T \mid q_{t+1} = S_j, \lambda) a_{ij} \\ &= \frac{\alpha_t(i) b_j(O_{t+1}) \beta_{t+1}(j) a_{ij}}{\sum_k \sum_l P(q_t = S_k, q_{t+1} = S_l, O \mid \lambda)} \\ &= \frac{\alpha_t(i) a_{ij} b_j(O_{t+1}) \beta_{t+1}(j)}{\sum_k \sum_l \alpha_t(k) a_{kl} b_l(O_{t+1}) \beta_{t+1}(l)} \end{split}$$

Maximum Likelihood

$$\xi_{t}(i,j) = P(q_{t} = S_{i}, q_{t+1} = S_{j} \mid O, \lambda)$$
  
$$\xi_{t}(i,j) = \frac{\alpha_{t}(i)a_{ij}b_{j}(O_{t+1})\beta_{t+1}(j)}{\sum_{k}\sum_{l}\alpha_{t}(k)a_{kl}b_{l}(O_{t+1})\beta_{t+1}(l)}$$

 We can calculate the probability of being in state S<sub>i</sub> at time t by marginalizing over the arc probabilities for all possible next states

$$\gamma_t(i) = \sum_{j=1}^N \xi_t(i,j)$$

- Maximum Likelihood
  - We can calculate the probability of being in state  $S_i$ at time t by marginalizing over the arc probabilities for all possible next states

$$\gamma_t(i) = \sum_{j=1}^N \xi_t(i,j)$$

- If the Markov model were not hidden but observable, both  $\gamma_t(i)$  and  $\xi_t(i, j)$  would be 0/1.
- In this case when they are not, we estimate them with posterior probabilities that give us *soft counts*

- Maximum Likelihood
  - We can calculate the probability of being in state  $S_i$ at time t by marginalizing over the arc probabilities for all possible next states

$$\gamma_t(i) = \sum_{j=1}^N \xi_t(i,j)$$

- In this case when they are not, we estimate them with posterior probabilities that give us *soft counts*
- We estimate posterior probabilities first (in the E-step) and calculate the parameters with these estimates (in the M-step).

- Baum-Welch algorithm
  - E-step computes  $\xi_t(i,j)$  and  $\gamma_t(i)$
  - M-step re-calculate  $\lambda$  given  $\xi_t(i, j)$  and  $\gamma_t(i)$
  - Until convergence
    - $P(O|\lambda)$  never decreases

Baum-Welch algorithm

 $- \text{E-step computes } \xi_{t}(i,j) \text{ and } \gamma_{t}(i)$   $E[z_{i}^{t}] = \gamma_{t}(i) \qquad z_{i}^{t} = \begin{cases} 1 & \text{if } q_{t} = S_{i} \\ 0 & \text{otherwise} \end{cases} z_{ij}^{t} = \begin{cases} 1 & \text{if } q_{t} = S_{i} \\ 0 & \text{otherwise} \end{cases}$   $E[z_{ij}^{t}] = \xi_{t}(i,j) \qquad z_{i}^{t} = \begin{cases} 1 & \text{if } q_{t} = S_{i} \\ 0 & \text{otherwise} \end{cases} z_{ij}^{t} = \begin{cases} 1 & \text{if } q_{t} = S_{i} \\ 0 & \text{otherwise} \end{cases}$   $\xi_{t}(i,j) = \frac{a_{t}(i)a_{ij}b_{j}(O_{t+1})\beta_{t+1}(j)}{\sum_{k}\sum_{l}a_{t}(k)a_{kl}b_{l}(O_{t+1})\beta_{t+1}(l)} \qquad \gamma_{t}(i) = \sum_{j=1}^{N}\xi_{t}(i,j)$ 

Baum-Welch algorithm

- E-step computes  $\xi_t(i,j)$  and  $\gamma_t(i)$ 

$$E[Z_{i}^{t}] = \gamma_{t}(1) \qquad z_{i}^{t} = \begin{cases} 1 & \text{if } q_{t} = S_{i} \\ 0 & \text{otherwise} \end{cases} \quad z_{ij}^{t} = \begin{cases} 1 & \text{if } q_{t} = S_{i} \text{ and } q_{t+1} = S_{j} \\ 0 & \text{otherwise} \end{cases}$$

$$E[Z_{ij}^{t}] = \xi_{t}(i,j) \qquad z_{i}^{t} = \begin{cases} 1 & \text{if } q_{t} = S_{i} \text{ and } q_{t+1} = S_{j} \\ 0 & \text{otherwise} \end{cases}$$

$$\xi_{t}(i,j) = \frac{a_{t}(i)a_{ij}b_{j}(O_{t+1})\beta_{t+1}(j)}{\sum_{k}\sum_{l}a_{t}(k)a_{kl}b_{l}(O_{t+1})\beta_{t+1}(l)} \qquad \gamma_{t}(i) = \sum_{j=1}^{N}\xi_{t}(i,j)$$

- M-step re-calculate  $\lambda$  given  $\xi_t(i, j)$  and  $\gamma_t(i)$ 

• Ratio of number of transitions from  $S_i$  to  $S_j$   $\hat{a}_{ij} = \frac{\sum_{t=1}^{T-1} \xi_t(i,j)}{\sum_{t=1}^{T-1} \gamma_t(i)}$ • Probability of observing  $v_m$  in  $S_j$   $\hat{b}_j(m) = \frac{\sum_{t=1}^{T} \gamma_t(j) \mathbf{1}(O_t = v_m)}{\sum_{t=1}^{T} \gamma_t(j)}$ 

Baum-Welch algorithm

– E-step computes  $\xi_t(i,j)$  and  $\gamma_t(i)$ 

$$\begin{aligned} E[z_i^t] &= \gamma_t(i) \\ E[z_i^t] &= \zeta_t(i,j) \end{aligned} z_i^t = \begin{cases} 1 & \text{if } q_t = S_i \\ 0 & \text{otherwise} \end{cases} z_{ij}^t = \begin{cases} 1 & \text{if } q_t = S_i \text{ and } q_{t+1} = S_j \\ 0 & \text{otherwise} \end{cases} \end{aligned}$$

$$\begin{aligned} E[z_i^t] &= \zeta_t(i,j) \\ \xi_t(i,j) &= \frac{a_t(i)a_{ij}b_j(O_{t+1})\beta_{t+1}(j)}{\sum_k \sum_l a_t(k)a_{kl}b_l(O_{t+1})\beta_{t+1}(l)} \end{aligned}$$

$$\begin{aligned} \gamma_t(i) &= \sum_{j=1}^N \xi_t(i,j) \end{aligned}$$

– M-step re-calculate  $\lambda$  given  $\xi_t(i, j)$  and  $\gamma_t(i)$ 

• For multiple observation sequences

$$\hat{a}_{ij} = \frac{\sum_{k=1}^{K} \sum_{t=1}^{T_k-1} \xi_t^k(i,j)}{\sum_{k=1}^{K} \sum_{t=1}^{T_k-1} \gamma_t^k(i)} \qquad \hat{\pi}_i = \frac{\sum_{k=1}^{K} \gamma_1^k(i)}{K}$$
$$\hat{b}_j(m) = \frac{\sum_{k=1}^{K} \sum_{t=1}^{T_k} \gamma_t^k(j) 1(O_t^k = v_m)}{\sum_{k=1}^{K} \sum_{t=1}^{T_k} \gamma_t^k(j)}$$

#### **Classification with HMMs**

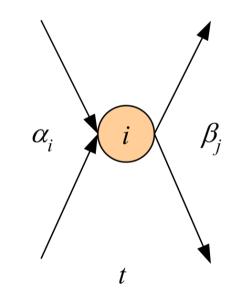
- We have a set of HMMs, each one modeling the sequences belonging to one class
  - For example, in spoken word recognition, examples of each word train a separate model,  $\lambda_i$ .
  - Given a new word utterance O to classify, all of the separate word models are evaluated to calculate  $P(O|\lambda_i)$ .
  - We then use Bayes' rule to get the posterior probabilities  $P(\lambda_i|O) = \frac{P(O|\lambda_i)P(\lambda_i)}{\sum_i P(O|\lambda_i)P(\lambda_i)}$

#### Exercise 1

• Prove the recursion expression for the forward-backward algorithm:

Recursion:

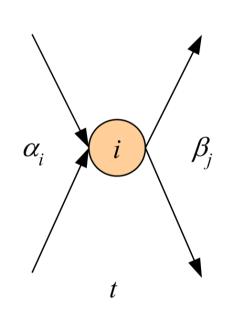
$$\beta_t(i) = \sum_{j=1}^N a_{ij} b_j(O_{t+1}) \beta_{t+1}(j)$$



#### Exercise 2

• Prove that for finding the state sequence

$$\gamma_{t}(i) \equiv P(q_{t} = S_{i} | O, \lambda)$$
$$= \frac{\alpha_{t}(i)\beta_{t}(i)}{\sum_{j=1}^{N} \alpha_{t}(j)\beta_{t}(j)}$$



#### Exercise 2

• Prove that for finding the state sequence

$$\gamma_{t}(i) \equiv P(q_{t} = S_{i} | O, \lambda)$$

$$= \frac{\alpha_{t}(i)\beta_{t}(i)}{\sum_{j=1}^{N} \alpha_{t}(j)\beta_{t}(j)}$$

$$\gamma_{t}(i) \equiv P(q_{t} = S_{i} | O, \lambda)$$

$$= \frac{P(O|q_{t} = S_{i}, \lambda)P(q_{t} = S_{i} | \lambda)}{P(O|\lambda)}$$

$$t$$

 $\beta_{j}$ 

#### Summary

Introduction Discrete Markov process Observable Markov model Hidden Markov model (HMM) Solving HMMs